

# Institutional Quality and Economic Performance in Nigeria: A New Evidence

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**Abstract:** The question of whether institutional quality is an important driver of growth has been the subject of a growing literature in both developed and developing economies across the globe. This study revisits this relationship in Nigeria from 1981Q1 to 2016Q4 and discusses the relevant policy implications for post Covid-19 Nigeria. The study adopted the ARDL approach which uses a bounds test approach based on unrestricted error correction model (UECM) to test for a long run relationship among the relevant variables. The findings indicate that institutional quality impacts negatively but insignificantly on growth in Nigeria, both at the aggregate and sectoral levels. However, initial output growth levels, capital and labour were found to be important drivers of growth in the country, while trade is growth-retarding. The study concludes that in this post Covid-19 era in Nigeria, there is need to improve the quality of socio-economic and political institutions in the country so that a more robust impact of these institutions can be felt in the economic performance of the country both at the aggregate and sectoral levels.

**Key words:** Institutional Quality; Economic Performance; ARDL Model; Nigeria

## I INTRODUCTION

The relationship between institutional quality and economic performance has been the subject of a growing literature in both developed and developing countries across the globe. This is because no

economy exists in a vacuum. There are certain institutional factors that play important roles in the growth of every economy. For instance, lack of political rights (i.e. freedom to participate in the political process), lack of respect for the rule of law and civil liberties (such as the rights to free expression, to organize or demonstrate, and to freedom of religion, education, travel, and other individual rights), poor government regulations, and high level of corruption have been blamed for the economic stagnation in most developing countries (Siba, 2008). North (1991) defined institutions as the rule of the game in a society or more formally, the humanly devised constraints that shape human interaction. Benyah (2010) explained institutional quality as the extent to which procedures by regulatory authorities foster investor protection and enhance greater access to funds for borrowers. Dysfunctional institutions generally render the economic environment unproductive and obstruct trade, thereby encumbering the growth objective of an economy (Anyanwu & Yaméogo, 2015)

Figure 1 (in appendix) reports the growth performance of the Nigerian economy. Panel A reports annual GDP growth while Panel B reports annual GDP per capita growth. Both panels aptly capture the poor performance of the Nigerian economy in recent years, especially the last recession of 2016 during which output plummeted and GDP per capita growth remained in the negative region. This raises concerns on the role of institutions in the Nigerian economic growth process. Thus, this paper examines the impact of

institutional quality on economic performance in Nigeria. However, unlike the extant literature which is based on the traditional institutional variables documented in the World Governance Indicators such as control of corruption, government effectiveness, regulatory quality, rule of law, political stability and absence of violence/terrorism, and voice and accountability, this paper derives an institutional quality variable based on the data from Freedom House, which monitors political freedom in various countries across the globe on annual basis.

## II AN OVERVIEW OF THE EMPIRICAL LITERATURE

Following the initial studies by Mauro (1995) and Knack and Keefer (1995), a growing empirical literature on the role of institutions in driving growth has emerged. Some aspects of this emerging literature show that institutions play important roles in Africa's growth performance. For instance, in a study aimed at establishing the factors contributing to Africa's growth performance, Mijiyawa (2013) found that government effectiveness was one of the key drivers of growth over the period 1995–2005. Similar results were obtained by Anyanwu (2014) who found that government effectiveness was an important driver of growth in Africa during the period 1996 - 2010. Furthermore, both Anyanwu and Yaméogo (2015) and Mijiyawa (2015) document that political instability, which is a major component of institutional quality, hinders foreign direct investment inflows into Africa, and thus impedes growth. Chikalipah (2017) documents that in SSA, strong institutional environment impacts positively and significantly on the performance of microfinance institutions, which constitute an important engine of growth in the sub-region. Other recent studies, such as Malikane and Chitambara (2017), show that strong democratic institutions promote economic growth in Southern African economies.

In West Africa, Iheonu et al. (2017) investigated the impact of institutional quality on economic performance from 1996 to 2015 using a panel of 12 West African economies. The study used the traditional measures of institutional quality such as control of corruption, government effectiveness, regulatory quality and rule of law, and adopted fixed effect, random effect and panel two-stage least square models. The results overwhelmingly showed support for positive and significant impact of institutions on economic performance. Other studies that have shown that institutional quality is an

important driver of growth in Africa, especially in West Africa include Ojapinwa (2017), Kebede and Takyi (2017), Amin (2013), and Kilishi et al. (2013).

In Nigeria, Dandume (2013) studied the relationship between institutions and growth performance using the ARDL approach. The findings indicate the existence of long run relationships between institutions and economic growth as well as a two-way causal relationship. Okoi, Okoi and Bassey (2015) also studied the relationship between institutional quality, macroeconomic policy, and economic development in Nigeria from 1995 to 2013, and found that contrary to Dandume (2013), the impact of domestic institution on Nigeria's development indices is insignificant. The findings of Okoi, Okoi and Bassey (2015) are however consistent with those of Iyoboyi and Pedro (2014), which also established that substantial amount of the changes in macroeconomic performance in Nigeria is not due to the quality of institutions. These conflicting findings in the extant literature for Nigeria means that a large scope still needs to be covered towards a more comprehensive understanding of the role of institutions in Nigeria.

Another recent study that has contributed towards a better understanding of the institutions-growth relationship in Nigeria is Olayungbo and Adediran (2017). This study investigated the effects of oil revenue and institutional quality on economic growth in Nigeria using annual data from 1984 to 2014 and the ARDL approach. Overall, the study established that institutional quality is important in explaining the relationship between oil revenue and economic growth in Nigeria. Akinkunmi (2017) also used the ARDL approach to establish that in the period 1960 to 2015, political stability and political freedom have impact negatively but insignificantly on economic growth in Nigeria.

The foregoing overview of the literature indicates that the debate on the role of institutions in Nigeria's economic growth has just begun. However, it must be stressed the extant literature for Nigeria is yet to derive institutional quality variable from Freedom House database, which monitors political freedom in each country on an annual basis. For each country, the Freedom House institutional quality dummy variable takes the value of 2 for the classification free, 1 for partly free, and 0 for not free. We find that the dummy index ranged between 0 and 2 for Nigeria. Studies like Alexiou et al. (2014) have used the Freedom House institutional quality data with great success. This study will exploit the Freedom House dataset to

investigate the role of institutional quality on economic performance in Nigeria.

### III METHODOLOGY

#### 3.1. Theoretical Framework:

The new growth theory indicates that growth of labour supply and growth of labour productivity are important factors in the economic growth process. Growth in labour productivity generally emanates from growth in human capital (i.e. accumulation of skills and knowledge), growth in investment (i.e. accumulation of physical capital), and technical progress (i.e. use of new and better production techniques). Following Tumwebaze and Ijjo (2015), we assume a Cobb-Douglas production function combining capital and labour with constant returns to scale so that aggregate output can be expressed as follows:

$$Y = AK^\alpha L^{1-\alpha} \quad (1)$$

where:  $Y$  = real economic output,  $A$  = technical progress,  $K$  = capital (proxied in this study by gross fixed capital formation in constant local currency), and  $L$  = labour (proxied by total adult population). The annual real GDP growth is obtained from equation (1) as:

$$y = a + \alpha k + (1 - \alpha)l \quad (2)$$

where:  $a$ ,  $y$ ,  $k$  and  $l$  denote the growth rates of  $A$ ,  $Y$ ,  $K$  and  $L$ , respectively. By assuming non-diminishing returns to the accumulation of both human capital and physical capital, the new growth theory is able to predict the long-term growth

$$PGDP_t = f(PGDP_{t-1}, K_t, L_t, FDI_t, TRADE_t, EXCH_t, INST_t) \quad (3)$$

where:  $PGDP_t$  is real per capita GDP growth;  $PGDP_{t-1}$  is the one lag of real per capita GDP growth; and  $t$  represents the time index. Furthermore, equation (3) is expressed as an ARDL model in its unrestricted ECM form as follows:

effects of institutional quality. Indeed, the extant literature has identified the channels through which institutional quality may impact on real per capita GDP growth. For instance, poor government regulations, lack of respect for the rule of law and civil liberties, lack of political rights and high level of corruption have been blamed for the economic stagnation in most developing countries, particularly the West African economies (Siba, 2008).

#### 3.2. Model Specification:

To model the growth effect of institutional quality in Nigeria, we extend the economic growth function in (2) by including standard variables in growth regressions that are specific to the Nigerian economy, such as foreign direct investment inflow ( $FDI_t$ ), trade openness measured as  $(Exports + Imports)/GDP$  (% of GDP) ( $TRADE_t$ ), nominal official exchange rate of local currency per U.S. dollar ( $EXCH_t$ ), and institutional quality ( $INST_t$ ). Even though these selected regressors are specific to the Nigerian economy, they are nonetheless consistent with some established studies in the literature such as Mankiw et al. (1992), Alexiou et al. (2014), and Tumwebaze and Ijjo (2015). The institutional quality variable used in this study is as defined in the last paragraph of Section 2. The entire data for this study is taken from the World Development Indicators published by the World Bank, except for the institutional quality variable, which was taken from Freedom House.

For the econometric analysis, we express the model for this study in its implicit form as follows:

$$\begin{aligned}
 \Delta PGDP_t = & \beta_0 + \beta_1 PGDP_{t-1} + \beta_2 K_{t-1} \\
 & + \beta_3 L_{t-1} + \beta_4 FDI_{t-1} \\
 & + \beta_5 TRADE_{t-1} + \beta_6 EXCH_{t-1} \\
 & + \beta_7 INST_{t-1} \\
 & + \sum_{i=1}^p \lambda_i \Delta PGDP_{t-i} \\
 & + \sum_{i=0}^q \alpha_i \Delta K_{t-i} + \sum_{i=0}^q \gamma_i \Delta L_{t-i} \\
 & + \sum_{i=0}^q m_i \Delta FDI_{t-i} \\
 & + \sum_{i=0}^q \psi_i \Delta TRADE_{t-i} \\
 & + \sum_{i=0}^q \delta_i \Delta EXCH_{t-i} \\
 & + \sum_{i=0}^q \phi_i \Delta INST_{t-i} \\
 & + \varepsilon_t \quad \dots \dots \dots \quad (4)
 \end{aligned}$$

where:  $\beta_0$  is the constant term;  $\varepsilon_t$  is the stochastic error term; and all the variables are logged prior to estimation. In what follows, we provide some explanations on how each variable in equation (4) is expected to influence real per capita GDP growth.

In line with economic theory, the parameters of labour, capital, FDI and trade are expected to have positive signs since these variables are expected to engender growth, though some studies have established that FDI may not enhance growth sometimes (Iheonu, 2016). Some studies provide empirical support for expansionary effects of devaluations, but the contractionary effects have become more prominent in recent large number of studies, though mixed results and insignificant effect have also been reported by few studies (Razzaque et al., 2017). Thus, the parameter of exchange rate may be positive or negative. Some recent empirical studies in the literature have provided evidence suggesting a positive relationship between institutions that promote economic freedom and economic performance (Iheonu et al., 2017; Wanjuu and Le Roux, 2017); while other recent empirical studies have also documented that weak and poor

institutional quality are growth retarding (Diop *et al.*, 2010; Ajide and Raheem, 2016). Hence, the coefficient of institutional quality is expected to be either positive or negative.

#### IV EMPIRICAL RESULTS AND DISCUSSION

This empirical investigation began with a review of the descriptive statistics of the variables as shown in Table 1. The variable names in Table 1 follow equation (4), except for the ones explained in the notes. We find from the mean, median, maximum and minimum values that the data shows the presence of no outliers. In addition, we find that the variables exhibit some variations exchange rate shows the highest level of variation while real GDP growth shows the least variation.

After the descriptive analysis, we examined the time series properties of the variables using the Phillips-Perron (PP) unit root test procedure. The results of this test are shown in Table 2. We find that all the variables are integrated of order one, indicating that they are not stationary at level. This suggests that a test for equilibrium relationship between the variables is important to guard against spurious regression. To do this, we adopted the bounds testing approach of Pesaran, Smith and Shin (2001, henceforth PSS). The results at both the aggregate and sectoral levels are shown in Table 3.

Table 1: Descriptive statistics of the variables

	AGRIC	CAPITAL	EXCH	FDI	IND	INST	LABOR	RGDP	SERV	TRADE
Mean	29.3830	28.9775	3.2840	21.1880	29.8917	0.8611	17.9500	12.3960	30.0134	3.8547
Median	29.1481	28.6692	3.9713	21.1718	29.7897	1.0000	17.9562	12.2639	29.7492	3.9773
Maximum	30.4566	30.0047	5.6431	22.9260	30.4542	2.0000	18.4190	12.8679	31.2393	4.4214
Minimum	28.4483	28.1788	-0.5179	18.9107	29.3274	0.0000	17.4936	12.0479	29.1608	3.0137
Std. Dev.	0.6591	0.5963	1.9363	1.1006	0.3411	0.5369	0.2749	0.2673	0.6866	0.3812
Skewness	0.2554	0.5372	-0.7297	-0.1301	0.0878	-0.1106	-0.0117	0.5695	0.5572	-0.6724
Kurtosis	1.5462	1.7091	2.1814	2.0934	1.6787	3.2050	1.7649	1.7174	1.8493	2.2967
Observations	144	144	144	144	144	144	144	144	144	144

Source: Authors. Notes: The variable names are as defined in equation (4), except for AGRIC, IND and SERV which denote agric, industrial and services value added in constant local currency, respectively. All the variables reported are logged, except for INST.

Table 2: Phillips-Perron (PP) Unit Root Test Results

Variable	PP Stat at Level	5% Critical Value	PP Stat at 1st Diff	5% Critical Value	Order of Integration
rgdp	-2.259122	-3.441552	-5.359429	-3.441777	I(1)
agric	-2.104674	-3.441552	-5.931677	-3.441777	I(1)
ind	-3.181416	-3.441552	-5.222118	-3.441777	I(1)
serv	-2.077678	-3.441552	-6.172591	-3.441777	I(1)
capital	-3.139946	-3.441552	-5.601017	-3.441777	I(1)
exch	-1.193364	-3.441552	-7.363059	-3.441777	I(1)
fdi	-2.968485	-3.441552	-8.461000	-3.441777	I(1)
labor	-2.802173	-3.441552	-4.234991	-3.441777	I(1)
trade	-1.628731	-3.441552	-7.804269	-3.441777	I(1)

Source: Authors.

Table 3: PSS Bounds Test Results

Dependent Variable	F-statistic	5% Lower Bound (I0 Bound)	5% Upper Bound (I1 Bound)	Decision
Rgdp	4.375912	2.45	3.61	Cointegration
Agric	5.095552	2.45	3.61	Cointegration
Ind	3.731531	2.45	3.61	Cointegration
Serv	6.583003	2.45	3.61	Cointegration

Source: Authors

The results in Table 3 indicate that a stable long-run relationship exists among the variables regardless of whether the aggregate log real GDP or the log sectoral value added is the dependent variable. This means that in all cases, the null hypothesis of no cointegration is rejected. The ARDL model for the aggregate output, using the real GDP growth as dependent variable, is reported in Table 4.

The results in Table 4 indicate that the long-run parameter for INST is given by  $-(0.38725)/1.40757 = -0.27512$ . However, the p-value of 0.0758 shows that this parameter is not statistically significant at the conventional 5% level. Table 4 also shows that in the short-run, INST impacts negatively on real GDP growth with a p-value of 0.9332, which is not statistically significant at the 5% level. Overall, we

find that the institutional quality variable impacts negatively but insignificantly on growth in Nigeria, both in the short-run and long-run. These findings are consistent with Okoi, Okoi and Bassey (2015) and Iyoboyi and Pedro (2014), which also found that the quality of institutions in Nigeria has not impacted significantly on macroeconomic performance. The findings are also consistent with Diop *et al.* (2010) and Ajide and Raheem (2016), which have also documented that weak and poor institutional quality is growth retarding. Indeed, these results are not surprising given that the Nigerian economy has been bedeviled by several issues relating to poor institutional quality such as endemic corruption, poor regulations, disregard for the rule of law, long years of military dictatorship/political instability, and so on.

Table 4 further shows that both in the long-run and short-run, the initial level of per capita income is an important driver of growth in Nigeria. This is consistent with the bulk of the extant literature such as Iheonu *et al.* (2017) and Levine and Renelt (1992). While Iheonu *et al.* (2017) established that initial per capita income is an important driver of growth in West Africa, Levine and Renelt (1992) found that the initial level of real GDP per capita is an important driver of GDP per capita growth in a

group of 101 countries. Other significant drivers of growth in the short-run are capital and labor, while trade is growth retarding.

To check if the foregoing results are robust to sectoral output growth in Nigeria, we estimated equation (4) separately using agric, industrial and services sectoral output growth as the dependent variables, respectively. The results are reported in the Appendix. Here, Panel 1 reports for the agric output growth equation, while Panels 2 and 3 report for industrial and services sectoral output growth equations, respectively. We find that even at the sectoral level, the role of institutional quality on output growth in Nigeria remained muted in the long-run throughout, except in the case of the industrial sector where it is significant at the 5% level. We also find that in the short-run, the impact of institutional quality on sectoral output growth remained unimportant all through. In fact, our results indicate that the impact of this variable is negative in all cases in the short-run, except for the services sector in which the impact is positive. Overall, we find that both at aggregate and sectoral levels, the impact of institutional quality on output growth is overwhelmingly negative but negligible.

Table 4: ARDL Estimation Results for Aggregate Output (Dependent Variable D(RGDP))

Variable	Coefficient	Std. Error	t-Statistic	Prob.
D(RGDP(-1))	0.44400	0.07353	6.0383	0.0000
D(RGDP(-2))	0.17972	0.07741	2.3216	0.0220
D(CAPITAL(-1))	0.05516	0.02327	2.3708	0.0194
D(EXCH)	-0.00682	0.00417	-1.6367	0.1044
D(FDI)	-0.00404	0.00653	-0.6188	0.5373
D(LABOR)	0.07754	0.03522	2.2014	0.0297
D(TRADE)	-0.06152	0.01590	-3.8679	0.0002
D(INST(-1))	-0.00056	0.00660	-0.0840	0.9332
RGDP(-1)	1.40757	0.07555	18.6321	0.0000
CAPITAL(-1)	-0.54870	0.44717	-1.2271	0.2223
EXCH(-1)	-0.18721	0.12184	-1.5366	0.1272
FDI(-1)	0.25865	0.12931	2.0002	0.0479
LABOR(-1)	2.12847	1.17799	1.8069	0.0734
TRADE(-1)	-0.23566	0.16128	-1.4612	0.1467
INST(-1)	0.38725	0.21610	1.7919	0.0758
Constant	-14.18246	12.82493	-1.1059	0.2711
Diagnosics				
R-squared	0.998322		F-statistic	2826.729
Adjusted R-squared	0.997969		Prob(F-statistic)	0.0000

Breusch-Godfrey Serial Correlation LM Test ( $\chi^2$ -stat):	7.166286 (0.0278)			
Breusch-Pagan-Godfrey Heteroskedasticity Test ( $\chi^2$ -stat):	50.33885 (0.0013)			

Source: Authors. Notes: Standard errors are Newey-West HAC standard errors. P-values are in parenthesis.

In the short-run, the results indicate that important drivers of growth include the initial level of output growth, capital (in the case of agric and services sectors) and labour (in the case of agric and industrial sectors), while trade is growth retarding all through. However, in the long-run, only labour exerts significant influence on output growth. These results substantially reflect the structure and dynamics of the Nigerian economy. For instance, the finding that trade is detrimental to growth in Nigeria is consistent with recent empirical evidence such as Iheonu *et al.* (2017) and Keho (2017), and this may be due to low trade volumes and negative trade balances recorded in Nigeria in the last two decades as well as the structural defects that characterize the economy such as poor transport infrastructure that hinders trade. Furthermore, the unimportant role of trade may also be explained by the fact that bilateral trade between Nigeria is more in terms of trade diversion than trade creation resulting from economic distortions that encourage illegal trade, which in turn significantly reduce bilateral trade between Nigeria and the rest of the world (Agbodji, 2008).

In terms of diagnostics, our results indicate the presence of the problems of autocorrelation and heteroskedasticity only in the aggregate equation reported in Table 4. However, we corrected for these problems using the Newey-West Heteroskedasticity and Autocorrelation Corrected (HAC) standard errors in order to restore the validity of the model. Besides, the ARDL approach adopted in this study corrects for any potential problem of weak endogeneity that usually bedevil growth equations through its lag structure. The F-Statistics indicate that all the models are statistically significant even at the 1% level.

## V. CONCLUSION AND POLICY RECOMMENDATIONS

The question of whether institutional quality is an important driver of growth has been the subject of a growing literature in both developed and developing economies across the globe. This study revisits this relationship in Nigeria from 1981Q1 to 2016Q4. However, unlike the extant literature which derived

institutional quality variables from the World Governance Indicators, this study built the institutional quality data from Freedom House database, which monitors political freedom across the globe. The study adopted the ARDL approach to modeling levels relationships recently advanced by Pesaran, Smith and Shin (2001). Overall, we find that the institutional quality variable impacts negatively but insignificantly on growth in Nigeria, both at the aggregate and sectoral levels. These results are robust both in the short-run and long-run. However, our results further indicate that initial output growth levels, capital and labour are important drivers of growth in Nigeria, while trade is growth-retarding.

The findings of this study have several policy implications. In this post Covid-19 era, policymakers in Nigeria are able to see that there is need to improve the quality of socio-economic and political institutions in the country so that a more robust impact of these institutions can be felt in the economic performance of the country. Specifically, the proposed improvements should target such areas as control of corruption, government effectiveness, regulatory quality, political stability, and respect for the rule of law and accountability. This means that policy reforms should be evolved to enhance these institutional frameworks in the country. Such reforms should be coordinated and escalated at all levels of government. From the findings of this study, policymakers in Nigeria are also able to see that higher levels of capital accumulation and human capital development are important for sustainable growth and development in the country. Thus, policy reforms in the country during this post Covid-19 period should support improved socio-political and economic environment that can attract higher levels of investment and enhance the quality of the labour force. This requires extensive structural transformations in order to make the economic environment more attractive to prospective investors. The proposed structural transformations can be achieved through the removal of bottlenecks to private and public investments, increased investment in basic infrastructure to drive productivity, increased government support to micro, small and medium scale enterprises, removal

of bottlenecks to transfer of land ownership and transparency in the enforcement of property rights. These will help Nigeria to maintain quality institutions that can attract more investments, increase trade and promote sustained economic growth and development even in this post Covid-19 era.

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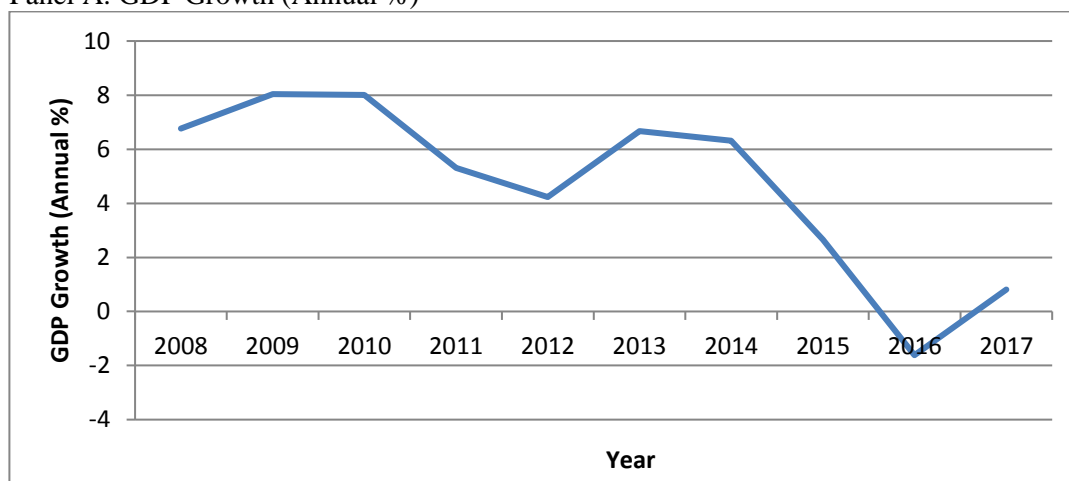
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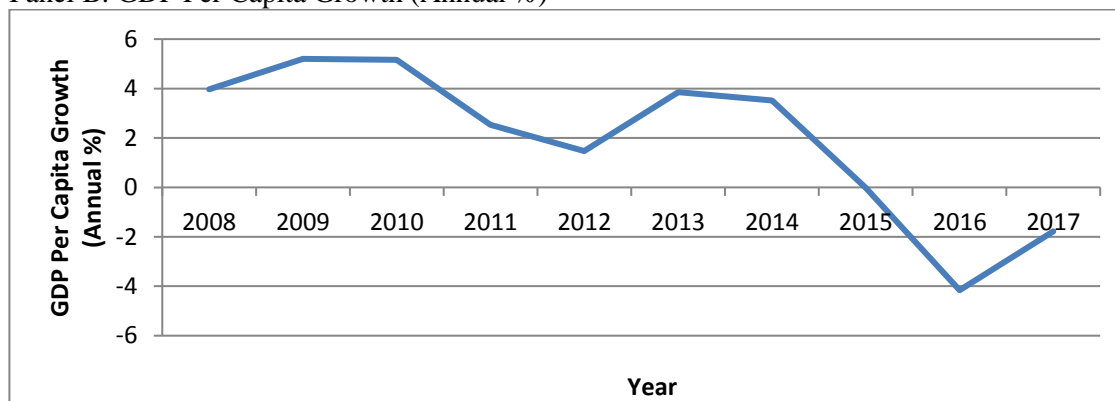
Appendix

Figure 1: GDP Growth Performance of Nigeria (2008 to 2017)

Panel A: GDP Growth (Annual %)



Panel B: GDP Per Capita Growth (Annual %)



Source: Authors, with data from World Development Indicators

Panel 1: ARDL Estimation Results for Aggregate Output (Dependent Variable D(AGRIC))

Variable	Coefficient	Std. Error	t-Statistic	Prob.
D(AGRIC(-1))	0.55029	0.07394	7.4421	0.0000
D(CAPITAL(-1))	0.14886	0.05749	2.5893	0.0110
D(EXCH)	0.00059	0.00588	0.1006	0.9200
D(FDI)	0.00424	0.00490	0.8648	0.3891
D(LABOR(-1))	19.79196	6.46971	3.0592	0.0028
D(TRADE)	-0.18925	0.03079	-6.1464	0.0000
D(INST(-1))	-0.00818	0.00956	-0.8554	0.3943
AGRIC(-1)	1.50188	0.07535	19.9330	0.0000

CAPITAL(-1)	-0.37827	0.25107	-1.5067	0.1349
EXCH(-1)	0.01221	0.12133	0.1007	0.9200
FDI(-1)	0.08758	0.08489	1.0316	0.3046
LABOR(-1)	2.68329	1.02219	2.6251	0.0099
TRADE(-1)	-0.39392	0.18724	-2.1038	0.0378
INST(-1)	0.10571	0.10186	1.0377	0.3017
CONSTANT	-6.70513	13.58068	-0.4937	0.6225
Diagnostics				
R-squared	0.999434		F-statistic	6042.534
Adjusted R-squared	0.999269		Prob(F-statistic)	0.0000
Breusch-Godfrey Serial Correlation LM Test ( $\chi^2$ -stat):	3.020856 (0.2208)			
Breusch-Pagan-Godfrey Heteroskedasticity Test ( $\chi^2$ -stat):	33.90587 (0.1688)			

Source: Authors. Notes: P-values are in parenthesis.

Panel 2: ARDL Estimation Results for Aggregate Output (Dependent Variable D(IND))

Variable	Coefficient	Std. Error	t-Statistic	Prob.
D(IND(-1))	0.78404	0.07380	10.6244	0.0000
D(IND(-2))	0.16935	0.07784	2.1757	0.0320
D(CAPITAL)	0.02324	0.02318	1.0025	0.3186
D(EXCH)	-0.00065	0.00383	-0.1708	0.8647
D(FDI)	-0.00437	0.00628	-0.6959	0.4881
D(LABOR)	20.88977	4.41877	4.7275	0.0000
D(TRADE(-1))	-0.10723	0.02764	-3.8789	0.0002
D(INST)	-0.00854	0.00486	-1.7561	0.0822
IND(-1)	1.70830	0.07860	21.7343	0.0000
CAPITAL(-1)	-0.05948	0.09017	-0.6596	0.5110
EXCH(-1)	-0.00863	0.05049	-0.1709	0.8647
FDI(-1)	0.06283	0.03550	1.7699	0.0798
LABOR(-1)	1.12545	0.40547	2.7756	0.0066
TRADE(-1)	0.04023	0.07939	0.5068	0.6134
INST(-1)	0.10437	0.04480	2.3298	0.0218
CONSTANT	9.84346	5.70840	1.7244	0.0878
Diagnostics				
R-squared	0.999234		F-statistic	3488.971
Adjusted R-squared	0.998947		Prob(F-statistic)	0.0000
Breusch-Godfrey Serial Correlation LM Test ( $\chi^2$ -stat):	2.949137 (0.2289)			
Breusch-Pagan-Godfrey Heteroskedasticity Test ( $\chi^2$ -stat):	30.13595 (0.6105)			

Source: Authors. Notes: P-values are in parenthesis.

Panel 3: ARDL Estimation Results for Aggregate Output (Dependent Variable D(SERV))

Variable	Coefficient	Std. Error	t-Statistic	Prob.
D(SERV(-1))	0.60944	0.07219	8.4420	0.0000
D(CAPITAL(-1))	0.05296	0.01932	2.7413	0.0070
D(EXCH)	0.00156	0.00344	0.4528	0.6515
D(FDI(-1))	-0.01161	0.00605	-1.9198	0.0572
D(LABOR)	5.88256	3.58740	1.6398	0.1036
D(TRADE)	-0.09100	0.01490	-6.1069	0.0000
D(INST)	0.00864	0.00466	1.8563	0.0658
SERV(-1)	1.60205	0.06946	23.0654	0.0000
CAPITAL(-1)	-0.85344	2.39159	-0.3568	0.7218
EXCH(-1)	0.21061	0.71074	0.2963	0.7675
FDI(-1)	0.01405	0.00609	2.3082	0.0226
LABOR(-1)	8.79336	3.54144	2.4830	0.0144
TRADE(-1)	-0.65118	1.15583	-0.5634	0.5742
INST(-1)	0.00864	0.00466	1.8563	0.0658
CONSTANT	25.74360	99.68120	0.2583	0.7966
Diagnostics				
R-squared	0.999763		F-statistic	30808.86
Adjusted R-squared	0.999731		Prob(F-statistic)	0.0000
Breusch-Godfrey Serial Correlation LM Test ( $\chi^2$ -stat):	2.948381 (0.229)			
Breusch-Pagan-Godfrey Heteroskedasticity Test ( $\chi^2$ -stat):	19.67419 (0.2353)			

Source: Authors. Notes: P-values are in parenthesis.