Unemployment rate and U.S. shadow economy: an analysis based on spline models

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Abstract—The paper aims to investigate the relationship between unemployment rate and shadow economy with USA data using spline models. The shadow economy is estimated as percentage of official GDP, using MIMIC model. The size of the shadow economy as % of official GDP is estimated using a MIMIC model with four causal variables (taxes on corporate income, contributions for government social insurance, unemployment rate and self-employment) and two indicators (index of real GDP and civilian labour force participation rate).The size of the shadow economy (SE) is estimated to be decreasing over the last two decades.

In order to evaluate the nature of the relationship between the two variables, we have estimated cubic B-spline, natural cubic B-spline and smoothing models.Using an F-test, we compare the smoothing spline to a global linear fit and the results indicate a sufficiently linear relationship. Finally, we have compared the local polynomial models with the spline model; the smoothing spline model closely matches the linearity between the size of the shadow economy and the unemployment rate. We extend the classical Okun law, in order to estimate the relationship between growth rate of official economy, unemployment rate and the size of the shadow economy. The results reveal a significant direct relationship between shadow economy and the unemployment rate and an indirect relation between shadow economy and growth of official sector.

***Keywords*—**shadow economy, unemployment rate, MIMIC models, spline models, Okun law, U.SA.

I. Introduction

The relationship between the shadow economy and the level of unemployment is one of major interest. People work in the shadow economy because of the increased cost that firms in the formal sector have to pay to hire a worker. The increased cost comes from the tax burden and government regulations on economic activities. In discussing the growth of the shadow economy, the empirical evidence suggests two important factors: (a) reduction in official working hours, (b) the influence of the unemployment rate.

Enste [16] points out that the reduction of the number of working hours below worker's preferences raises the quantity of hours worked in the shadow economy. Early retirement also increases the quantity of hours worked in the shadow economy.

In Italy, Bertola and Garibaldi [6] present the case that an increase in payroll taxation can have effect on the supply of labour and the size of the shadow economy. An increase in tax and social security burdens not only reduces official employment but tends to increase the shadow labour force. This is because an increase in payroll tax can influence the decision to participate in official employment. Also, Boeri and Garibaldi [7] show a strong positive correlation between average unemployment rate and average shadow employment across 20 Italian regions during 1995-1999.

The paper examines the possible relationship between unemployment rate and the size of the shadow economy using a nonparametric analysis based on spline models. Also, a reexamination of the classical Okun's law is provided in the paper, showing the relationship between unemployment and official economy in the presence of shadow economy.

II. Estimating the Size of the U.S. Shadow Economy

II.1. Data and Methodology

II.1.1. Data issues

The variables used in the estimation are defined in appendix A. The data series are quarterly, seasonally adjusted covering the period 1980:Q1 to 2009:Q2.

The series in levels or differences have been tested for unit roots using the Augmented-Dickey Fuller (ADF) test and PP tests. All the data has been differentiated for the achievement of the stationarity (appendix, unit root analysis). While all the variables have been identified like integrated on first order, the latent variable is estimated in the same transformation of independent variables (first difference).

II.1.2 Methodology

The size of the U.S. shadow economy is estimated as % of official GDP using a particular type of structural equations models-MIMIC model.

The MIMIC model- Multiple Indicators and Multiple Causes model (MIMIC model), allows to consider the SE as a

"latent" variable linked, on the one hand, to a number of observable indicators (reflecting changes in the size of the SE) and on the other, to a set of observed causal variables, which are regarded as some of the most important determinants of the unreported economic activity [10].

The model is composed by two sorts of equations, the structural one and the measurement equations system. The equation that captures the relationships among the latent variable (η) and the causes (X_q) is named "structural model" and the equations that links indicators (Y_p) with the latent variable (non-observed economy) is called the "measurement model".

A MIMIC model of the hidden economy is formulated mathematically as follows:

$$Y = \lambda\eta + \varepsilon \quad (1)$$

$$\eta = \gamma' X + \xi \quad (2)$$

where:

η is the scalar latent variable(the size of shadow economy);

$Y' = (Y_1,....Y_p)$ is the vector of indicators of the latent variable;

$X' = (X_1,...X_q)$ is the vector of causes of η ;

$\lambda_{(p\times 1)}$ and $\gamma_{(q\times 1)}$ vectors of parameters;

$\varepsilon_{(p\times 1)}$ and $\xi_{(q\times 1)}$ vectors of scalar random errors;

The $\varepsilon' s$ and ξ are assumed to be mutually uncorrelated.

Substituting (2) into (1), the MIMIC model can be written as:

$$\mathrm{Y} = \Pi\mathrm{X} + \mathrm{z} \quad (3)$$

where: $\Pi = \lambda\gamma^{'}, z = \lambda\xi + \varepsilon$.

The estimation of (1) and (2) requires a normalization of the parameters in (1), and a convenient way to achieve this is to constrain one element of λ to some pre-assigned value ([17]-[18]).

The possible causes of shadow economy considered in the model are: tax burden decomposed into personal current taxes (X_1), taxes on production and imports(X_2), taxes on corporate income(X_3), contributions for government social insurance(X_4) and government unemployment insurance(X_5), unemployment rate(X_6), self-employment in civilian labour force (X_7), government employment in civilian labour force (X_8) called bureaucracy index. The indicator variables incorporated in the model are: real gross domestic product index (Y_1), currency ratio M_1/M_2 (Y_2) and civilian labour force participation rate (Y_3).

The variables used into the estimation of the shadow economy are also quarterly and seasonally adjusted covering the period 1980-2009. All the data has been differentiated for the achievement of the stationarity.

In order to estimate the MIMIC model, by Maximum Likelihood, using the LISREL 8.8 package, we normalized the coefficient of the index of real GDP ($\lambda_1 = -1$) to sufficiently identify the model. This indicates an inverse relationship between the official and shadow economy.

In order to identify the best model, we have started with MIMIC model 8-1-3 and we have removed the variables which have not structural parameters statistically significant.

A detailed description and implementation of the MIMIC model for the USA shadow economy is provided in [15].

II.2. Empirical results

In order to estimate the size of the shadow economy, we have identified the best model as MIMIC 4-1-2 with four causal variables (taxes on corporate income, contributions for government social insurance, unemployment rate and self-employment) and two indicators (index of real GDP and civilian labour force participation rate).

Taking into account the reference variable ($Y_1, \frac{\mathrm{Re}\,al\,GDP_t}{\mathrm{Re}\,al\,GDP_{1990}}$) the shadow economy is scaled up to a value in 1990, the base year, and we build an average of several estimates from this year for the U.S.A. shadow economy (table I).

The index of changes of the shadow economy (η) in United States measured as percentage of GDP in the 1990 is linked to the index of changes of real GDP as follow:

Measurement Equation: $$\frac{GDP_t - GDP_{t-1}}{GDP_{1990}} = \frac{\tilde{\eta}_t - \tilde{\eta}_{t-1}}{GDP_{1990}} \quad (4)$$

I: Estimates of the size of U.S.A. shadow economy (1990)

Author	***Method***	***Size of Shadow Economy***
Johnson et. Al(1998)	Currency Demand Approach	13.9%
Lacko(1999)	Physical Input(Electricity)	10.5%
Schneider and Enste(2000)	Currency Demand Approach	7.5%*
Mean 1990		***10.6%***

*means for 1990-1993

The estimates of the structural model are used to obtain an ordinal time series index for latent variable (shadow economy):

Structural Equation:

$$\frac{\tilde{\eta}_t}{GDP_{1990}} = -0.24\Delta X_{3t} + 3.00\Delta X_{4t} + 1.49\Delta X_{6t} + 1.01\Delta X_{7t} \quad (5)$$

The index is scaled to take up to a value of 10.6% in 1990 and further transformed from changes respect to the GDP in the 1990 to the shadow economy as ratio of current GDP:

$$\frac{\widetilde{\eta}_t}{GDP_{1990}} \times \frac{\eta^*_{1990}}{GDP_{1990}} \times \frac{GDP_{1990}}{\widetilde{\eta}_{1990}} \times \frac{GDP_{1990}}{GDP_t} = \frac{\hat{\eta}_t}{GDP_t} \quad (6)$$

I. $\frac{\widetilde{\eta}_t}{GDP_{1990}}$ is the index of shadow economy calculated by eq.(2);

II. $\frac{\eta^*_{1990}}{GDP_{1990}} = 10.6\%$ is the exogenous estimate of shadow economy;

III. $\frac{\widetilde{\eta}_{1990}}{GDP_{1990}}$ is the value of index estimated by (5);

IV. $\frac{GDP_{1990}}{GDP_t}$ is to convert the index of changes respect to base year in shadow economy respect to current GDP;

V. $\frac{\hat{\eta}_t}{GDP_t}$ is the estimated shadow economy as a percentage of official GDP.

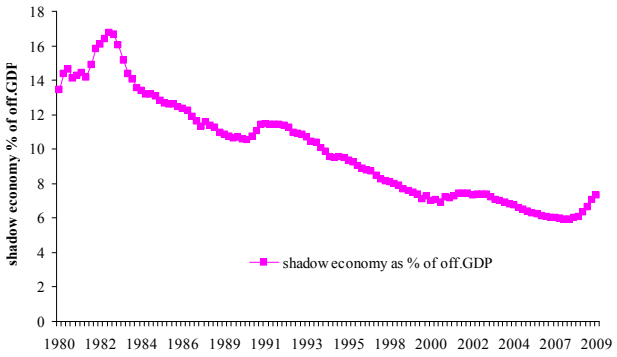


Fig. 1. The size of the shadow economy in United States as % of official GDP

The shadow economy measured as percentage of official GDP records the value of 13.41% in the first trimester of 1980 and follows an ascendant trend reaching the value of 16.77% in the last trimester of 1982. At the beginning of 1983, the dimension of USA shadow economy begins to decrease in intensity, recording the average value of 6% of GDP at the end of 2007. For the last two year 2008 and 2009, the size of the unreported economy it increases slowly, achieving the value of 7.3% in the second quarter of 2009.

The results are not far from the last empirical studies for USA ([16], [29]).Schneider estimates in his last study, the size of USA shadow economy as % of GDP, at the level of 7.9% in 2005, respectively 8% in 2006.

III. THERE IS A LINK BETWEEN SHADOW ECONOMY AND UNEMPLOYMENT RATE IN THE CASE OF UNITED STATES?

In many empirical studies, is has been found that tax burden is the biggest causes of shadow economy. Also the size of shadow economy is influenced by the level of unemployment. An increase in unemployment rates reduces the proportion of workers employed in the formal sector' this leads to higher labor participation rates in the informal sector.

The graphical evolution of the shadow economy versus unemployment rate reveal the existence of a strong positive relationship between the two variables, quantified by a value of about 0.80 of correlation coefficient.

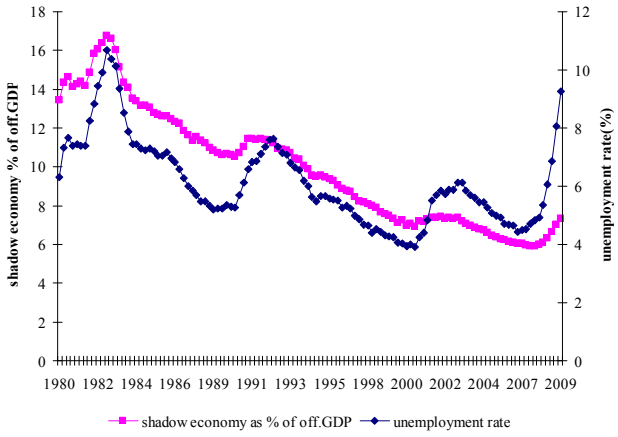


Fig.2. Shadow economy vs. Unemployment rate in United States

Giles ([17], [18]) states that the effect of unemployment on the shadow economy is ambiguous (i.e. both positive and negative). An increase in the number of unemployed increases the number of people who work in the black economy because they have more time. On the other hand, an increase in unemployment implies a decrease in the shadow economy. This is because the unemployment is negatively related to the growth of the official economy (Okun's law) and the shadow economy tends to rise with the growth of the official economy

III.1. A Nonparametric Analysis of the Relationship between Unemployment Rate and the Size of the Shadow Economy Using Spline Models

The paper aim to investigate the relationship between shadow economy estimated using the MIMIC model and unemployment rate using a nonparametric analysis based on spline models. The unemployment rate is expressed in %, taken from U.S. Bureau of Statistics, Labour Force Statistics from Current Population Survey.

Instead of assuming that we know the functional form for a regression model, a better alternative is to estimate the functional form from the data, replacing global estimates with

local estimates. In the terms of local estimation, the statistical dependency between two variables is described not with a single parameter such as a mean or a slope coefficient, but with a series of local estimates.

Like local polynomial regression (LPR), spline smoothers are another nonparametric technique used with scatterplots. In any spline model, it must be selected the number of knots and the knot placement [23].

Standard practice is to place knots at evenly spaced intervals in the data. But the question of how to select the number of knots remains and has an important effect on the Spline fit. One method is to use a visual trial. Four knots is the standard starting point. If the fit appears rough, knots are added. If the fit appears overly nonlinear, knots are subtracted. The second method is to use Akaike Information Criterion to select the number of knots. The optimal number of knots is returned by the lowest AIC value.

Thus far, LPR estimates have revealed a linear dependency between the size of the shadow economy and the unemployment rate. It will we interesting to investigate the nature of the relationship between the two variables, using both cubic B-splines and natural cubic B-splines to estimate the nonparametric fit.

For the both spline models, it has been used 4 knots it we will evaluate whether this is the optimal number of knots, using Akaike Criterion.

Analyzing the graphics of both functions, there is a little difference between cubic B-splines and the natural cubic B-splines.

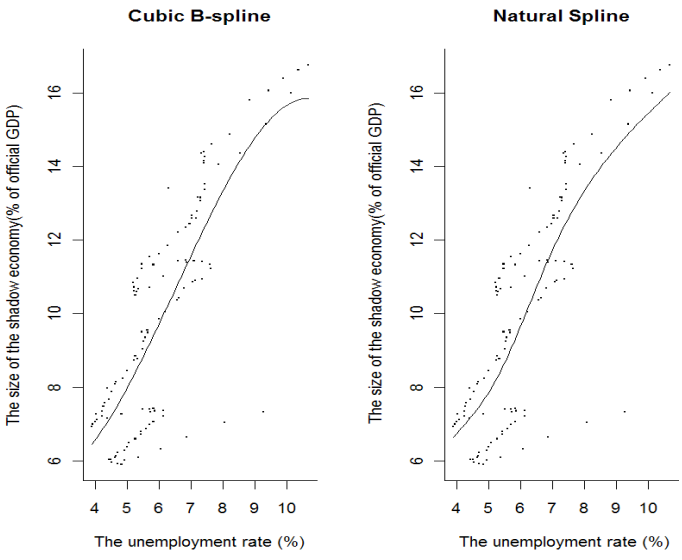


Fig.2. Cubic B-spline and natural spline fit to size of the shadow economy (% of off.GDP)

In order to select the number of knots, we use for both models the Akaike Information Criterion (AIC); the optimal number of knots is returned by the lowest AIC value. For the both spline models, it has been estimated several models with 2-9 knots.

II: AIC values for differing number of knots

	Natural Spline	**Cubic Spline**
2 knots	477.0573	478.7849
3 knots	478.1010	478.7849
4 knots	477.0685	478.7849
5 knots	479.0302	478.7093
6 knots	480.0504	479.2769
7 knots	***475.1998***	477.6035
8 knots	477.5902	479.2186
9 knots	477.2197	***475.646***

Analysing the values of Akaike Information Criterion for several knots we observe that the optimal number of knots for the Natural Spline taking into account the lowest AIC value is 7 knots, while for the cubic Spline the optimal number of knots is 9.

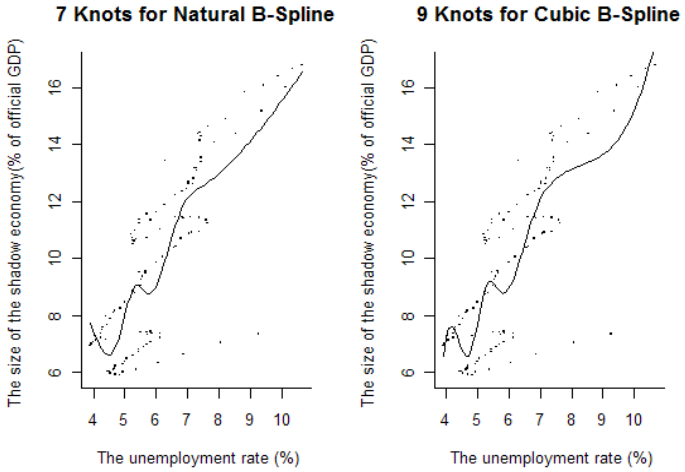


Fig.3. The choice of the optimal number of knots for the both models

In order to estimate the statistical relationship between two variables, both splines and local polynomial regression can provide such an estimate with few assumptions about functional form. A common criticism of the both methods is that it is easy to have a surfeit of local parameters, which produces overly nonlinear estimates that overfit data [23].

Penalized splines are a nonparametric regression technique that minimizes the possibility of overfitting. Smoothing splines operate with penalized estimation, placing a penalty on the number of local parameters used to estimate the nonparametric fit.

Like linear regression models, the spline estimate $\hat{f}$ minimises the sum of squares between y and the nonparametric estimate, $f(x_i)$:

$$SS(f) = \sum_{i=1}^{n} [y - f(x)]^2 \qquad (7)$$

The main problem is that the estimate of f that minimises (7) use too many parameters. The penalised estimation solution is to attach a penalty for the number of parameters used to estimate f: $\lambda\int_{x_1}^{x_n}[f''(x)]^2dx$, named roughness penalty that have two components: λ, the smoothing parameter and the second, the integrated squared second derivative of $f(x)$.

Further, the spline estimate become:

$$SS(f,\lambda)=\sum_{i=1}^{n}[y-f(x)]^2+\lambda\int_{x_1}^{x_n}[f''(x)]^2dx \qquad (8)$$

While small values of λ will interpolate the data and large values returns a least squares fit, intermediate values does not offer an interpretable effect on the amount of the smoothing applied to the data. It is proposed a transformation of λ into an approximation of the degrees of freedom; by selecting the degrees of freedom, it is chosen the number of effective local parameters used in the spline estimate.

The penalized splines named also "smoothing splines" differ from the standard splines by the fact that the number of knots have little influence over how smooth the fit is since the value of λ controls now the quality of the fit.

In order to see how different degrees of freedom (2, 4, 8 and 12) affect the fit, it has been estimate the relationship between the size of the shadow economy and the unemployment rate using smoothing splines.

The results reveal that the fit with 2 degrees of freedom is identical to a linear regression; for the model with 4 and 8 degrees of freedom we have the same pattern of linearity found with other spline fits. For the fit with 12 degrees of freedom, we have considerable variability, caused by too many parameters and we can conclude that the data are over fitted.

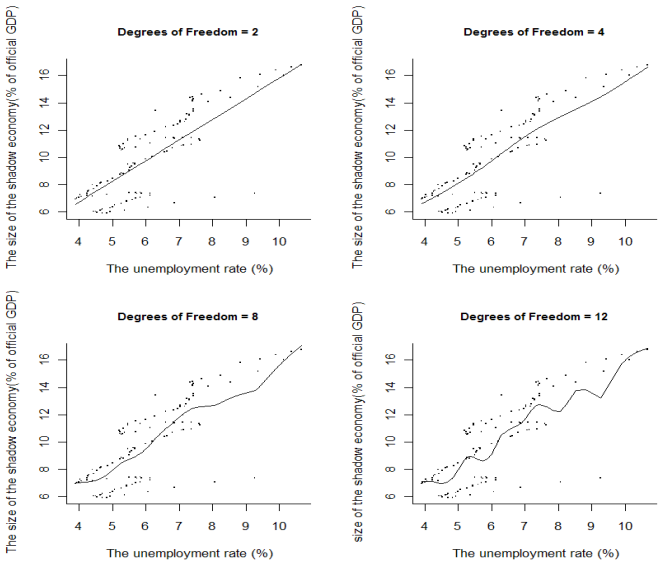


Fig.4. Smoothing spline fit to shadow economy data

In order to test hypothesis about the nature of the relationship between the size of the shadow economy and the unemployment rate, we compare the smoothing spline model to a model with only a constant to test whether the effect of the unemployment rate is significantly different from zero.

If RSS_1 and RSS_2 are the residual sum of squares from a restricted model and the spline model respectively, the

$$F-test=\frac{(RSS_1-RSS_2)/(df_{res2}-df_{res1})}{RSS_2/(n-df_{res2})}\approx F_{df_{res2}-df_{res1},n-df_{res2}} \qquad (9)$$

Applying the F-test, we find that the relationship between the two variables is highly significant as the test statistic is 72.718 on 3 and 115 degrees of freedom (p= 2.2e-16). We also test the spline model against a global linear fit, and the value of F-test of 1.9597 on 2 and 114 degrees of freedom is not statistically significant(p=0.1476). The results of the test indicate that the relationship between the size of the shadow economy and the unemployment rate is sufficiently linear and the global linear fit is adequate.

Finally, we provide a comparison of the nonparametric regression models (fig.5).

The first two nonparametric models are the loess and lowess smoothers. In Dobre, Alexandru [9] we have estimated the both models identifying the optimal value of span at the value of 0.4. Between the two LPR smoothers, the lowess estimate provides a better fit of the data.

In the lower left panel is the natural cubic B-spline with 7 knots, chosen by the AIC values. This model displays noticeable undersmoothing of the estimate. Finally, we estimate a smoothing spline using 4 degrees of freedom selected through visual trial. The smoothing spline closely matches the linearity between the size of the shadow economy and the unemployment rate.

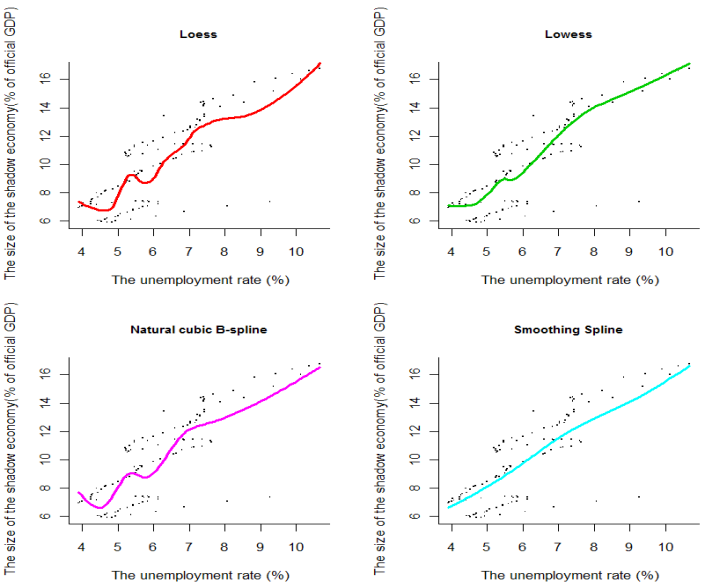


Fig. 5. Comparison of smoother fits

III. 2. A re-examination of Okun's law in presence of shadow economy

The Okun's law relates decreases in the unemployment rate to increases in output growth. We want to test if the shadow economy has any significant effect on this empirical evidence. We go on the hypothesis that a lower growth rate of official GDP from potential output is associated with higher deviations of the unemployment rate from its "natural" level. The increase in unemployment leads to an increase in the number of laborers who work in the unofficial labour market.

In fig.1(appendix), we present the significant statistical relationships among growth rate of official GDP, changes in unemployment rate and growth of shadow economy for the case of United States covering the period 1980-2009.

The estimates obtained based on the standard relation given by Okun's law are presented in the following table:

$$g_t^Y = \alpha_0 \Delta u_t + \varepsilon_t \quad (10)$$

where:

$g_t^Y = (g_t^{off} - \bar{g}_{(80-09)}^Y)$ indicates the difference of growth rate of the official gross domestic product (g_t^{off}) from it average calculated over the period 1970 to 2008;

$g_t^\eta = (g_t^{shad} - \bar{g}_{(80-09)}^\eta)$ indicates the difference of shadow economy(g_t^{shad}) from it average calculated over the period 1980 to 2009, Δu_t id the first difference of unemployment rate, ε_t are residuals i.i.d.

III. Estimation output of regression:

$$g_t^Y = \alpha_0 \Delta u_t + \varepsilon_t$$

Dependent Variable: G_GROWTH
Method: Least Squares
Date: 07/15/10 Time: 20:20
Sample: 1980Q2 2009Q2
Included observations: 117

	Coefficient	Std. Error	t-Statistic	Prob.
DU	-1.754908	0.166910	-10.51413	0.0000
R-squared	0.487965	Mean dependent var		0.000171
Adjusted R-squared	0.487965	S.D. dependent var		0.852275
S.E. of regression	0.609859	Akaike info criterion		1.857333
Sum squared resid	43.14368	Schwarz criterion		1.880941
Log likelihood	-107.6540	Hannan-Quinn criter.		1.866918
Durbin-Watson stat	1.810181			

The estimates show an inverse relationship between changes in unemployment and the growth rate of official output. Furthermore, we use a modified version of Okun's law by including the shadow economy:

$$g_t^Y = \alpha_1 \Delta u_t + \beta g_t^\eta + \varepsilon_t \quad (11)$$

IV. Estimation output of regression:

$$g_t^Y = \alpha_1 \Delta u_t + \beta g_t^\eta + \varepsilon_t$$

Dependent Variable: G_GROWTH
Method: Least Squares
Date: 07/15/10 Time: 20:23
Sample: 1980Q2 2009Q2
Included observations: 117

	Coefficient	Std. Error	t-Statistic	Prob.
DU	-0.696164	0.279337	-2.492196	0.0141
G_ETA	-1.686008	0.370776	-4.547237	0.0000
R-squared	0.565999	Mean dependent var		0.000171
Adjusted R-squared	0.562225	S.D. dependent var		0.852275
S.E. of regression	0.563904	Akaike info criterion		1.709080
Sum squared resid	36.56854	Schwarz criterion		1.756296
Log likelihood	-97.98115	Hannan-Quinn criter.		1.728249
Durbin-Watson stat	1.508212			

The econometric results reveal that we have a significant negative relationship on the one hand, between the growth rate of official economy and the level of unemployment, that confirm the Okun's law, and on the other hand, between the growth rate of official output and the size of the shadow economy. We deduce therefore, that shadow economy tends to cushion the effects of changes in unemployment on the official GDP.

In order to investigate the impact of shadow economy on the unemployment rate, we develop a structural relationship, taking into account also the growth rate of official GDP:

$$g_t^{shad} = \gamma g_t^{off} + \lambda \Delta u_t + \varepsilon_t \quad (12)$$

where:

(g_t^{off}) is the first difference of annual growth rate of the official gross domestic product;

g_t^{shad} is the first difference of the shadow economy;

Δu_t is the first difference of unemployment rate; ε_t residuals;

V. Estimation output of regression:

$$g_t^{shad} = c + \gamma g_t^{off} + \lambda \Delta u_t + \varepsilon_t$$

Dependent Variable: G_SHAD
Method: Least Squares
Date: 07/15/10 Time: 20:32
Sample: 1980Q2 2009Q2
Included observations: 117

	Coefficient	Std. Error	t-Statistic	Prob.
C	0.001980	0.019761	0.100205	0.9204
G_OFF	-0.088836	0.019845	-4.476588	0.0000
DU	0.474043	0.050005	9.479919	0.0000

R-squared	0.744924	Mean dependent var	-0.052154
Adjusted R-squared	0.740449	S.D. dependent var	0.255095
S.E. of regression	0.129961	Akaike info criterion	-1.217857
Sum squared resid	1.925446	Schwarz criterion	-1.147032
Log likelihood	74.24465	Hannan-Quinn criter.	-1.189103
F-statistic	166.4625	Durbin-Watson stat	1.345612
Prob(F-statistic)	0.000000		

The parameter γ of the equation shows an inverse relationship between the growth of the official economy (g_t^{off}) and growth of the shadow economy (g_t^{shad}). On the other-hand, the parameter λ shows a direct relationship between changes in unemployment and the growth of the shadow economy.

The coefficients are statistically significant (prob.<5%) and the degree of determination in the model is high, 75% of the variation of shadow economy is explained by the two exogenous variables unemployment rate and growth rate of official GDP.

Our estimations show that the presence of the shadow economy acts as a buffer as it absorbs some of the unemployed workers from the official economy into the shadow economy.

IV. CONCLUSIONS

The main goal of the paper is to investigate the nature of the relationship between unemployment rate and the size of the shadow economy of the USA data using spline models.The shadow economy is estimated as percentage of official GDP, using MIMIC model. The results show that the size of the shadow economy varies from thirteen to seventeen percent between 1980 and 1983 and then decreases steadily up to 7 percent of official GDP in 2009.

We investigate the nature of the relationship between the two variables, using cubic B-splines and natural cubic B-splines to estimate the nonparametric fit. The graphics of both spline models reveals a little difference between the two functions. Using an F-test, we compare the smoothing spline model to a model with only a constant, and we conclude that unemployment rate has a statistically significant effect on the size of the U.S.A. shadow economy.

We also test the spline model against a global linear fit and the results indicate that the relationship between the size of the shadow economy and the unemployment rate is sufficiently linear and the global linear fit is adequate.

Finally, we have compared the local polynomial regression models (loess and lowess) estimated in Dobre, Alexandru[15] with spline models (natural cubic B-spline and smoothing spline). From the four types of models that we have applied, the smoothing spline model closely matches the linearity between the size of the shadow economy and the unemployment rate.

We extend the classical Okun's law, in order to estimate the relationship between growth rate of official economy, unemployment rate and the size of the shadow economy.

The results reveal a significant direct relationship between shadow economy and the unemployment rate and an indirect relation between shadow economy and growth of official sector.

Moreover, we can conclude that employment in the shadow economy constitutes a form of labor market transition between or rather from unemployment back into formal employment.

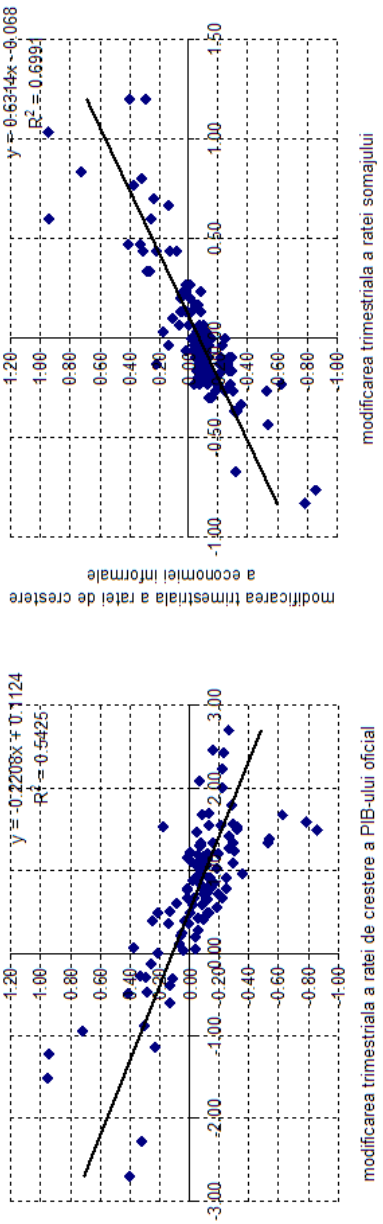
Appendix. A. Unit-root analysis

The data sources are: Bureau of Economic Analysis (BEA), Bureau of Labor Statistics Data (BLS) and Federal Reserve Banks.

	CAUSES	Source	Unit root analysis		Level				First Difference				
					ADF	lag	PP	lag	ADF	lag	PP	lag	
X_1	Personal current taxes/GDP	BEA	I(1)	T&C	-2.474	4	-2.11	7	-13.83*	0	-13.42*	7	$\Delta(X_1)$
				C	-2.493	4	-2.00	6	-13.79*	0	-13.40*	7	
				None	-0.881	1	-6.761	6	-13.79*	0	-13.40*	7	
X_2	Taxes on production and imports/GDP	BEA	I(1)	T&C	-3.543	0	-3.813	6	-11.24*	0	-11.28*	2	$\Delta(X_2)$
				C	-2.922	0	-3.286	6	-11.27*	0	-11.31*	2	
				None	0.289	0	0.255	5	-11.32*	0	-11.37*	2	
X_3	Taxes on corporate income/GDP	BEA	I(1)	T&C	-4.19*	3	-3.45	6	-10.93*	0	-10.93*	4	$\Delta(X_3)$
				C	-4.14*	3	-3.44	6	-10.98*	0	-10.98*	4	
				None	-1.18	0	-1.13	5	-4.14*	2	-11.01*	4	
X_4	Contributions for government social insurance/GDP	BEA	I(1)	T&C	-2.32	6	-2.01	10	-5.99*	5	-18.57*	8	$\Delta(X_4)$
				C	-2.73	4	-3.40	10	-4.36*	3	-14.03*	9	
				None	0.62	4	1.08	10	-4.35*	3	-13.75*	9	
X_5	Government unemployment insurance	BEA	I(1)	T&C	-2.63	2	-1.76	4	-4.44*	2	-6.49*	3	$\Delta(X_5)$
				C	-3.22	2	-2.07	4	-4.00*	1	-6.36*	3	
				None	-0.94	2	-0.38	4	-4.33*	2	-6.37*	3	
X_6	Unemployment rate	BLS	I(1)	T&C	-1.03	1	-1.41	6	-4.40*	0	-4.69*	3	$\Delta(X_6)$
				C	-2.14	1	-1.69	6	-4.17*	0	-4.52*	3	
				None	-0.22	1	0.03	7	-4.17*	0	-4.53*	3	
X_7	Self-employment/Civilian labour force	BLS	I(1)	T&C	-2.44	0	-2.18	4	-9.68*	1	-12.79*	13	$\Delta(X_7)$
				C	-0.90	0	-0.51	5	-11.03*	0	-11.44*	8	
				None	-0.71	0	-0.91	9	-11.05*	0	-11.41*	8	
X_8	Index of bureaucracy	BLS	I(1)	T&C	-2.69	0	-2.67	3	-12.73*	0	-12.38*	7	$\Delta(X_8)$
				C	-2.88	0	-2.94	3	-12.01*	0	-11.99*	7	
				None	-2.27	0	-2.10	5	-11.37*	0	-11.68*	7	
	INDICATORS												
Y_1	M_1 / M_2	Federal Reserve Banks	I(1)	T&C	-2.12	2	-1.43	8	-3.02	1	-6.59*	7	$\Delta(Y_1)$
				C	-1.69	2	-0.90	8	-3.03*	1	-6.51*	7	
				None	-0.59	2	-0.78	8	-3.02*	1	-6.48*	7	
Y_2	Index of Real GDP[1]	BEA	I(1)	T&C	-1.71	2	-2.35	4	-5.43*	1	-8.71*	4	$\Delta(Y_2)$
				C	1.14	2	2.03	5	-5.26*	1	-8.44*	4	
				None	4.63	2	9.68	5	-2.39*	1	-4.45*	6	
Y_3	Civilian labor force participation rate	BLS	I(1)	T&C	-0.47	2	-0.66	1	-10.29*	1	-10.59*	0	$\Delta(Y_3)$
				C	-2.03	0	-2.08	3	-5.98*	2	-10.08*	4	
				None	1.12	0	1.15	4	-5.81*	2	-10.01*	5	

[1] Real Gross Domestic Product, Chained Dollars. Billions of chained (2500) dollars. Seasonally adjusted at annual rates/ $\mathrm{Re}al\ GDP_{1990-Q1}$

Note:
T&C represents the most general model with a drift and trend; C is the model with a drift and without trend; None is the most restricted model without a drift and trend. Numbers in brackets are lag lengths used in ADF test (as determined by SCH set to maximum 12) to remove serial correlation in the residuals. When using PP test, numbers in brackets represent Newey-West Bandwith (as determined by Bartlett-Kernel). Both in ADF and PP tests, unit root tests were performed from the most general to the least specific model by eliminating trend and intercept across the models (Katircioglu, 2009). *, ** and *** denote rejection of the null hypothesis at the 1%, 5% and 10% levels respectively. Tests for unit roots have been carried out in E-VIEWS 6.0.



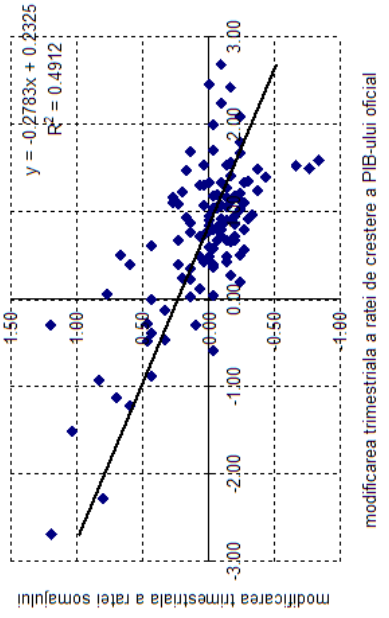


Fig.1.Growth of official GDP, Changes in unemployment and Growth of Shadow Economy

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*** R 2.9.1 software